

# Pre-processing for Gibbs Random Fields

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Aidan Boland, Nial Friel



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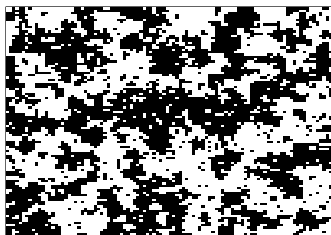


# Outline

- Short overview of Gibbs random fields.
- Outline current techniques and their pitfalls.
- Introduce an approach to improving the computational cost of analysis.

# What is a Gibbs random field?

- Type of graphical model.

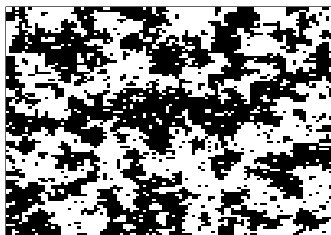


- Set of nodes with corresponding random variables ( $y$ ).
- $y_i \in [-1, 1]$

- Our interest is the posterior  
posterior  $\propto$  likelihood  $\times$  prior  
 $\pi(\theta|y) \propto f(y|\theta)\pi(\theta)$
- However the likelihood is intractable for GRF's

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# Intractability of GRF's

- The likelihood function is intractable

$$f(y|\theta) = \frac{q(y|\theta)}{Z(\theta)} = \frac{\exp(\theta^T s(y))}{Z(\theta)},$$

- $\theta$  tells us about the properties of the GRF.
  - $s(y)$  are summary statistics for the graph.
- $Z(\theta) = \sum_y \exp(\theta^T s(y))$ .
- $Z(\theta)$  is intractable for all but trivial graphs.
    - $16 \times 16$  binary graph,  $2^{16 \times 16} = 1.15 \times 10^{77}$  calculations.
    - $30 \times 30$  binary graph,  $2^{30 \times 30} = 8.453 \times 10^{270}$  calculations!

# Examples of intractable methods

- Finance: Stochastic volatility model.
- Ecology: Spatial models for biodiversity.
- Genetics: Phylogenetic trees.
- Sociology: Social Networks
- Epidemiology: Stochastic models for disease transmission
- Machine Learning: Deep architecture

# Current Methods

- Want to sample from the posterior  $\pi(\theta|y)$
- Can use MCMC sampling.
- Metropolis-Hastings.
  - Propose to move to  $\theta' \sim h(\cdot|\theta)$ ,
  - Accept move with probability  $\alpha$ ,

$$\begin{aligned}\alpha &= \min \left( 1, \frac{f(y|\theta')}{f(y|\theta)} \frac{\pi(\theta')}{\pi(\theta)} \frac{h(\theta|\theta')}{h(\theta'|\theta)} \right) \\ &= \min \left( 1, \frac{q(y|\theta')}{q(y|\theta)} \frac{Z(\theta)}{Z(\theta')} \frac{\pi(\theta')h(\theta'|\theta)}{\pi(\theta)h(\theta|\theta')} \right)\end{aligned}$$

## Current Methods (Exchange Algorithm)

- Sample instead from the augmented distribution

$$\pi(\theta', y', \theta | y) \propto f(y|\theta)\pi(\theta)h(\theta'|\theta)f(y'|\theta')$$

- Sample  $y' \sim f(\cdot|\theta')$
- Acceptance ratio becomes

$$\begin{aligned} \alpha &= \min \left( 1, \frac{q(y|\theta')}{q(y|\theta)} \frac{\cancel{Z(\theta)}}{\cancel{Z(\theta')}} \frac{\pi(\theta')h(\theta'|\theta)}{\pi(\theta)h(\theta|\theta')} \frac{q(y'|\theta)}{q(y'|\theta')} \frac{\cancel{Z(\theta)}}{\cancel{Z(\theta')}} \right) \\ &= \min \left( 1, \frac{q(y|\theta')}{q(y|\theta)} \frac{\pi(\theta')h(\theta'|\theta)}{\pi(\theta)h(\theta|\theta')} \frac{q(y'|\theta)}{q(y'|\theta')} \right) \end{aligned}$$



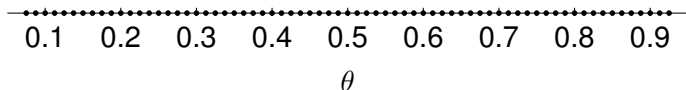
# Current Methods (ABC)

- Approximate Bayesian Computation
- Only requires the ability to sample from the likelihood
  - Sample  $y' \sim f(\cdot|\theta')$
  - Accept  $\theta'$  if  $\rho(s(y) - s(y')) < \varepsilon$
  - Where  $\rho$  is some distance function.
- ABC MCMC
- Use acceptance ratio,

$$\alpha = \min \left( 1, \frac{\pi(\theta')h(\theta'|\theta)}{\pi(\theta)h(\theta|\theta')} \mathbb{1}[\rho(s(y) - s(y')) < \varepsilon] \right)$$

# Pre-processing

- Computationally expensive to sample from  $f(y|\theta)$
- Prior to algorithm, sample  $y'$  at fixed set of  $\theta$  values ( $\theta_{pre}$ ).

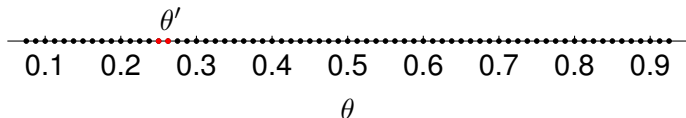


- When a sample is required, can interpolate from these pre sampled values.

$$s_{\theta'}(y') \sim N\left(\mathbb{E}[s_{\theta'_{pre}}(y)], \text{Var}(s_{\theta'_{pre}}(y))\right).$$

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# Estimating the normalising ratio

- Importance sampling,

$$\frac{Z(\theta)}{Z(\theta')} = \mathbb{E}_{f(y'|\theta')} \frac{q(y'|\theta)}{q(y'|\theta')} \approx \frac{1}{N} \sum_{i=1}^N \frac{q(y'_i|\theta)}{q(y'_i|\theta')}$$

- Numerical integration,

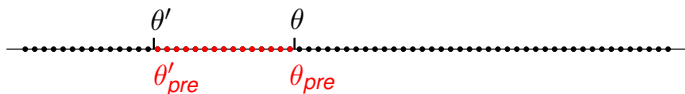
$$\log \left[ \frac{Z(\theta)}{Z(\theta')} \right] = \int_{\theta'}^{\theta} \mathbb{E}_{y|\theta}[\mathbf{s}(y)] \approx \sum_{i=a}^b \left\{ (\theta_{i+1} - \theta_i) \frac{\mathbb{E}_{\theta_i}[\mathbf{s}(y)] - \mathbb{E}_{\theta_{i+1}}[\mathbf{s}(y)]}{2} - \frac{(\theta_{i+1} - \theta_i)^2}{12} (V_{\theta_i}(\mathbf{s}(y)) - V_{\theta_{i+1}}(\mathbf{s}(y))) \right\}$$

# Estimating the normalising ratio

- Propose  $\theta' \sim h(\cdot|\theta)$ .
- Find nearest pre computed values,
  - $\theta \approx \theta_{pre}$
  - $\theta' \approx \theta'_{pre}$ .

- Approximate  $\frac{Z(\theta)}{Z(\theta')} \approx \frac{\widehat{Z(\theta_{pre})}}{\widehat{Z(\theta'_{pre})}}$ .

- Where  $\frac{\widehat{Z(\theta_{pre})}}{\widehat{Z(\theta'_{pre})}}$  is either the importance sampling or numerical integration estimate.



# Correction

- We can use the pre computed values to improve the accuracy of the estimate.

$$\frac{Z(\theta_{pre})}{Z(\theta)} \approx \frac{\widehat{Z(\theta_{pre})}}{Z(\theta)} = \frac{1}{N} \sum_{i=1}^N \frac{q(y'_i | \theta_{pre})}{q(y'_i | \theta)}$$

- Where  $y'_i$  is the  $i^{\text{th}}$  pre computed sample at  $\theta_{pre}$ .

- Finally  $\frac{Z(\theta)}{Z(\theta')} \approx \frac{\widehat{Z(\theta'_{pre})}}{Z(\theta')} \times \frac{\widehat{\widehat{Z(\theta_{pre})}}}{\widehat{Z(\theta'_{pre})}} \times \frac{\widehat{Z(\theta)}}{\widehat{Z(\theta_{pre})}}$

## Theoretical Guarantee

- Want  $|\alpha(\theta, \theta') - \hat{\alpha}(\theta, \theta', y')| \leq \delta(\theta, \theta')$
- Letting  $\hat{R}_i = \frac{\widehat{Z(\theta_{i+1})}}{\widehat{Z(\theta_i)}}$ .
- We get the following bound

$$\mathbb{P} \left( \left| \prod_{i=1}^M \hat{R}_i - \prod_{i=1}^M \mathbb{E} [\hat{R}_i] \right| \leq \gamma \right) \geq 1 - \varepsilon$$

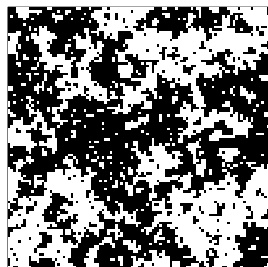
with  $\gamma = \alpha(1 + (M - 1)MK^{M-1})$

and  $\alpha = (b - a) \sqrt{\frac{1}{2N} \log \left( \frac{2M}{\varepsilon} \right)}$ .

- where
  - M is the number of pre processed  $\theta$ 's between  $\theta$  and  $\theta'$ .
  - N is the number of samples at each  $\theta_{pre}$
  - K is the upper bound on the ratio of likelihoods at any two consecutive values of  $\theta_{pre}$ .

# Ising

- Single parameter model.
- Defined on a rectangular lattice.
- Models spatial distribution of binary variables  $(-1, 1)$ .
- $f(y|\theta) = \frac{1}{Z(\theta)} \exp(\theta s(y))$ 
  - $s(y) = \sum_{j=1}^N \sum_{i \sim j} y_i y_j$
  - $i \sim j$  denotes that  $i$  and  $j$  are neighbours.
  - $\theta$  determines the degree of association between neighbours.



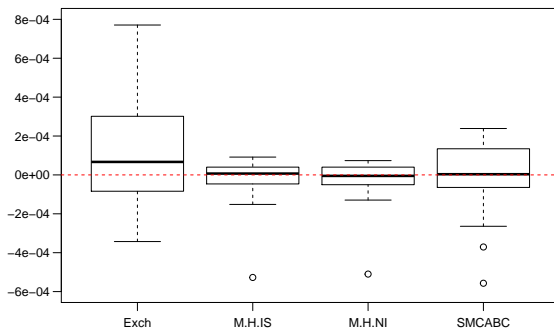
Ising example



# Simulation Study

- Simulated  $24 \times 80 \times 80$  Ising graphs.
- Ran an exchange algorithm for 24 hours to use as a 'ground truth'.
- Pre computation took 16 minutes per graph.

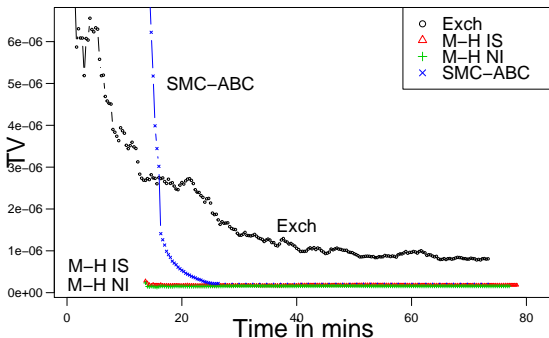
**Mean bias (first 20 mins)**



# Simulation Study

- Simulated 24 graphs of size  $80 \times 80$ .
- Ran an exchange algorithm for 24 hours to use as a 'ground truth'.
- Pre computation took 16 minutes per graph.

### Total Variation Graph 22



# END

- Any questions?

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An Roinn Proif, Fostair agus Nuálaíochta  
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